The cost-efficiency carbon pricing puzzle*

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Abstract

Any global temperature target must be translated into an intertemporal carbon budget and its associated cost-efficient carbon price schedule. Following Hotelling’s rule, the growth rate of this price should be equal to the interest rate. It is therefore a puzzle that cost-efficiency IAM models yield carbon prices that increase at an average real growth rate around 7% per year. This carbon pricing puzzle suggests that their abatement trajectories are not intertemporally optimized, probably because of the political acceptability constraint. Using an intertemporal asset pricing approach, I show that the uncertainties surrounding economic growth and future abatement technologies can partially solve this puzzle. I calibrate a simple two-period version of the model by introducing infrequent macroeconomic catastrophes à la Barro in order to fit the model with observed assets pricing in the economy. I show that marginal abatement costs and aggregate consumption are positively correlated, implying a positive carbon risk premium and an efficient growth rate of expected carbon prices larger than the interest rate. From this numerical exercise, I recommend a growth rate of expected carbon price around 3.75% per year (plus inflation). I also show that the rigid carbon budget approach to cost-efficiency carbon pricing implies a large uncertainty surrounding the future carbon prices that support this constraint. In this model, green investors are compensated for this risk by a large risk premium embedded in the growth rate of expected carbon prices, not by a collar on future carbon prices as often recommended.

Keywords: Carbon budget, Hotelling’s rule, consumption-based CAPM, climate finance.

JEL codes: Q54, D81, G12.

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1 Introduction

At the occasion of the COP-21 in Paris in 2015, the target of 2°C (and possibly 1.5°C) has been confirmed. This "Paris Agreement" must be translated into an intertemporal carbon budget constraint for the next few decades. Determining the optimal timing to consume this carbon budget is a problem isomorphic to the Hotelling’s problem of extracting a non-renewable resource (Hotelling (1931), Chakravorty et al. (2006), Chakravorty et al. (2008), Schubert (2008)). Under this cost-efficiency approach, abating one ton of CO\(_2\) today is a perfect substitute to abating one ton of CO\(_2\) in the future.\(^1\) Frontloading the abatement effort is an investment that has a single cost and a single benefit that are respectively equal to the present and future marginal abatement costs (MAC). If the climate policy is decentralized through carbon pricing (a carbon tax or a market for bankable permits), this corresponds to the present and future carbon prices. Along the optimal abatement path, this marginal investment should have a zero net present value. This is possible only if the growth rate of (expected) carbon price is equal to the (risk-adjusted) discount rate.\(^2\) This Hotelling’s rule applied to climate change is simple and transparent.\(^3\) The ambition of the climate target or the emergence of low-cost abatement technologies should influence the initial carbon price, but not its growth rate over time.

In most climate models, there is no uncertainty and green technological progresses are known in advance. In that case, the growth rate of carbon prices should be the socially desirable rate to discount risk-free investments. In an economy with efficient financial markets, it should be the interest rate. It is then a puzzle that most of these models generate carbon prices whose real growth rate is much larger than the interest rate. At the occasion of the publication of its 5th report in 2014, the IPCC established a database of several climate models. If one limits the analysis to the 767 calibrations of these models that estimate a world carbon prices for years 2020 and 2050, they yield an average annual growth rate of 7.04% for real carbon prices between these two dates, with a median of 5.70% and a standard deviation of 4.48%. The histogram of the annualized real growth rate of carbon price is described in Figure 1. Given the low interest rates in our economies, this cannot be intertemporally efficient. I refer to this observation as the "carbon pricing puzzle" of cost-efficient IAM models.

In the absence of any credibility problem, the decentralization of the allocation of an intertemporal carbon budget over 30 years, one should emit the corresponding number of permits that could be used at any time during the period, i.e., the system should allow full permits banking. Under certainty, the growth rate of equilibrium carbon price should equal the interest rate. The violation of this property for the models of the IPCC is a positive version of the carbon pricing puzzle. In reality, these models explore second-best climate policies in

\(^1\)In this introduction, I assume that the rate of natural decay of CO\(_2\) in the atmosphere is zero.

\(^2\)In this introduction, I ignore the natural decay of GHG in the atmosphere. In reality, the growth rate of carbon price should be equal to the sum of the discount rate and of the rate of natural decay.

\(^3\)It is specific to the cost-efficient approach implied by the Paris Agreement. It does not need to hold in the cost-benefit approach used for example by Nordhaus (2018). For example, using a 3% discount rate, the U.S. administration published a scientific report (IAWG (2016)) based on a cost-benefit approach that recommends a price of 42 dollars (of 2007) per ton of CO\(_2\) in 2020, growing to 69 dollars (of 2007) in 2050. This yields a real growth rate of 1.65% per year. Because the carbon concentration in the atmosphere will continue to grow over time under the optimal mitigation strategy, carbon prices will grow in parallel to the MAC, assuming a convex damage function.
Mean = 7.04%
Median = 5.70%
St. dev. = 4.48%

Figure 1: Histogram of the real growth rate (in % per year) of carbon prices between 2020 and 2050 from 767 calibrations of IAM models contained in the IPCC database (https://tntcat.iiasa.ac.at/AR5DB).

which the intertemporal allocation of the carbon budget is not optimized. Rather, these models characterize carbon price schedules that are compatible with exogenously determined carbon emission targets at different dates. These 'Representative Concentration Pathways' (RCP) are predetermined by the IPCC. The large growth rate of carbon prices suggests that the waiting game of international climate politics has infected the IPCC.

In the United States, the Climate Leadership Council (CLC) has sponsored an 'Economists' Statement on Carbon Dividends' in early 2019. The CLC supports a carbon price at 40 USD\textsubscript{2017} growing at 5% per year above inflation. In France, a recent public report (Quinet (2019)) has recommended a carbon price growing from 69 EUR\textsubscript{2019} to 775 EUR\textsubscript{2019} between 2020 and 2050, yielding a real growth rate of 8% per annum. In the United Kingdom, the official public carbon values grow from 14 GBP\textsubscript{2018} in 2020 to 81 GBP\textsubscript{2018} in 2030 (BEIS (2019)), implying a growth rate of 16% per annum. Because of their large growth rates of prices, these recommendations raise two difficulties. From a normative point of view, a reallocation of the abatement efforts to the present would increase intertemporal welfare under the same total carbon budget constraint. Even if these carbon prices would be implemented in the future, these plans illustrate the waiting game in the global fight against climate change. From a positive viewpoint, suppose that carbon pricing be implemented through a market for bankable emission permits. Under the carbon prices described above, investors would be willing to buy as many permits as possible today to sell them in the future, whereas firms would be willing to save as many permits as possible. This would imply an excess demand for permits. Therefore, these price schedules cannot support an equilibrium. Permit banking should therefore be prohibited and the problem of the intertemporal allocation of mitigation efforts goes back to the political arena, a bad news given the difficulty for politicians to commit on a long-term environmental policy.
This puzzle suggests that neither the RCPs of the IPCC nor the national NDCs of the Paris Agreement are intertemporally optimized, and that this inefficiency will induce difficulties to decentralize these quantitative targets through carbon pricing. However, the puzzle is based on the premise that the frontloading strategy is risk-free and that the future price of carbon is certain. In this paper, I recognize that these key assumptions are unrealistic, and I explore the impact of uncertainty on the socially efficient growth rate of real carbon prices. From a positive viewpoint, I predict the growth rate of expected carbon prices if the intertemporal carbon budget is decentralized through a market for permits with full banking.

The abatement models using a cost-efficiency approach and a carbon budget rely on strong assumptions about the evolution of the abatement cost function during the next few decades (Pindyck (2013)). Obviously, technologically optimistic models allow for low prices and efforts in the short run by anticipation of the emergence of these low-cost mitigation technologies. But in reality, technological evolutions are very hard to predict. If they do not materialize, one will have to drastically increase carbon prices to satisfy the intertemporal carbon budget. Nobody really knows today what will be the cost of abatement associated to wind or solar energy in 30 years. And deep uncertainties surround future electricity storage technologies and nuclear fusion for example. The extraordinary large uncertainty surrounding the emergence of economically viable renewable systems of energy is an inherent dimension of the energy transition. Similarly, these abatement models are generally based on a deterministic growth of total factor productivity. Recognizing the uncertainty surrounding the growth of TFP in the long run should also be taken into account to determine the carbon price schedule. If economic growth is larger than expected, more abatement efforts will have to be implemented to compensate for the larger emissions and this will also require a larger carbon price. One should also recognize that scientific progresses about climate sciences could induce us to revise the carbon budget downwards or upwards. As in the "quantity" approach proposed by Weitzman (1974) under uncertainty, I assume that the carbon budget is not sensitive to changes in the marginal abatement costs. This implies that modelers around the world face large challenges to implement this cost-efficiency approach to carbon pricing. In this paper, I focus the analysis on how uncertainty affects the efficient rate of growth of the carbon price.

The uncertainty of future carbon prices and MACs should affect the optimal timing of climate efforts and the carbon pricing system that support it. Common wisdom suggests that this uncertainty should induce us to implement strong immediate actions to reduce emissions. This suggests that uncertainty should push carbon prices up in the short run, thereby allowing for a reduction of the growth rate of carbon price to satisfy the carbon budget. I show in this paper that the opposite is true: Uncertainty tends to increase the efficient growth rate of carbon price, thereby allowing for a smaller carbon price today. The theoretical question raised here is about how to adapt the Hotelling’s rule to uncertainty. There has been a few attempts to answer this question in the late XXth century. For example, Pindyck (1978, 1980) explores the optimal extraction strategy of risk-neutral owners of a nonrenewable resource when exploration is possible or when the stock of this resource and the demand for it are unknown. This analysis is useful to examine a resource-rich country that is unable or unwilling to make this asset financially liquid, but it is not directly relevant in the context of the carbon budget problem. Indeed, households, investors and firms that will bear the mitigation risk will also bear all other statistically-linked risks in the economy. Our approach is closer to Gaudet and Howitt (1989), Gaudet and Khadr (1991) and Slade and Thille (1997) who examined the case of stochastic processes for economic growth and extraction costs in the context of a non-renewable resource.
normative version of the Consumption-based Capital Asset Pricing Model (CCAPM, Breeden (1979), Lucas (1978) and Rubinstein (1976)) tells us how the uncertainty affecting the future benefits of abatement frontloading should affect their efficient pricing, and thus the socially desirable abatement strategy. From a more positive viewpoint, it also tells us how early adopters of green technologies should be compensated for the contribution of their early green investments on the aggregate risk that they bear at equilibrium.

A risk premium should therefore be added to the interest rate to determine the trend of growth of carbon price in order to take account of the impact of the climate policy on the macroeconomic risk. Suppose for example that, along the optimal path, marginal abatement costs are negatively correlated with aggregate consumption. Because the MAC is the future benefit of abatement frontloading, fighting climate change early has the extra benefit to hedge the macro risk in that context. Because of this negative CCAPM beta of early mitigation efforts, one should discount the future benefit of this early investment, i.e., the future MAC, at a rate lower than the risk-free rate to determine the current price of carbon. This means at the same time a larger current price of carbon, and a growth rate of the expected carbon price smaller than the risk-free rate. From a positive point of view, this carbon pricing system is compatible with an equilibrium, as investors in green technologies will have an expected rate of return smaller than the interest rate, just because such green investments hedge their global portfolio risk. On the contrary, if MAC and aggregate consumption correlate positively, i.e., if the climate beta is positive, the risk premium will be positive, the current price of carbon will be smaller, and the growth rate of expected carbon price will be larger than the interest rate. This policy provides the right price signal for private investors in renewables technologies to take account of the impact of their decisions on social welfare, as is the case on efficient financial markets for other investment projects.

It remains thus to characterize the determinants of this carbon beta. To do this, I develop a two-period "act-then-learn" model in which the dynamically optimal mitigation strategy is endogenously determined under uncertainty about the abatement cost function, economic growth and carbon budget. I characterize the impact of these sources of uncertainty on the optimal growth rate of expected carbon price, and I realistically calibrate this model. Most integrated assessment models which allow for uncertainty do that by using a "learn-then-act" methodology. Under that approach, it is assumed that the modeler observes the realization of the vector of uncertain parameters before optimizing the climate policy under certainty. The mean value of the conditionally optimal policies across all possible realizations of this vector is then recommended as the optimal policy under uncertainty. This method is not satisfactory because it ignores the timing of the resolution of the uncertainty, and therefore the role of precaution that is inherent to our world. It also produces carbon prices that are not coherent with the system of assets prices in the economy. By using stochastic dynamic optimization and backward induction, I more realistically determine the optimal climate policy in the first period before the resolution of the uncertainty taking place in the second period. I solve the classical asset pricing puzzles (Mehra and Prescott (1985), Weil (1989) and Kocherlakota (1996)) of the CCAPM by introducing catastrophes in the growth process, as suggested by Barro (2006). In this framework, I show that the beta

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5Dietz et al. (2018) examined the risk profile of carbon prices using the cost-benefit analysis of the DICE model. In this alternative approach, the key determinant of the climate beta is the income-elasticity of climate damages.

6I also examine a model in which the asset pricing puzzles are resolved by using the long run risk approach.
of abatement frontloading is the income-elasticity of MACs. Multiplying this beta by the equilibrium aggregate risk premium tells us by how much the growth rate of expected carbon price should differ from the equilibrium interest rate. I show that the sign of this carbon beta is generally ambiguous, with different sources of uncertainty pushing the climate beta in opposite directions. However, a realistic calibration of the two-period model suggests a positive climate beta. This means that it is socially desirable to implement a climate strategy with a growth rate of expected carbon price that is larger than the interest rate, thereby allowing to start with a relatively low carbon price today. Thus, this analysis justifies using a discount rate for green technologies and planning for a growth rate of expected carbon prices that are larger than the interest rate. It could thus help solving the carbon pricing puzzle. However, the efficient growth rate of carbon prices is around 3.5%, which is much smaller than the 7.04% observed on average in the database of models of the IPCC. The bottom line of my analysis remains that the RCPs of the IPCC inefficiently allocate abatement efforts over time. The same final concentration of GHG in the atmosphere could be obtained with a smaller impact on intergenerational welfare by abating more today, and abating less in the future.

A possible explanation of the carbon pricing puzzle is based on the existence of political constraints related to the social acceptability of climate policies around the world in the short run. Following Gollier and Tirole (2015) for example, these constraints are typically at play to postpone climate efforts to the future, a phenomenon of procrastination that could explain why the above-mentioned models support a low current carbon price and a large growth rate of this price. This raises the question of the credibility of long-term climate commitments. Laffont and Tirole (1996) take this question seriously by proposing a commitment device based on forward financial contracts. Harstad (2019) justifies strategic investments and investment subsidies in technologies that are strategic complements to future green investments when the social planner faces a time-consistency problem from hyperbolic discounting.

In the next section, I assume that the optimal abatement strategy under the carbon budget is known, and I characterize the properties of the carbon pricing system that supports this social optimum, assuming an exogenous statistical relation between MAC and aggregate consumption. Section 3 is devoted to a simple two-period model in which the price of carbon in the first period must be determined under uncertainty about economic growth, green innovation and carbon budget. The carbon beta is determined endogenously in this section. In Section 4, I calibrate this model.

2 The efficient growth rate of expected carbon price

In this section, we examine a simple dynamic model of exogenous growth and technological uncertainty. Consider an economy with a fixed carbon budget. Suppose that this carbon budget has been allocated intertemporally in an optimal way. We determine the properties of the schedule of carbon prices that supports this optimum. In the spirit of the CCAPM, suppose that the economy has a representative agent whose rate of pure preference for the present is $\rho$. The von Neumann-Morgenstern utility function $u$ of the representative agent is increasing and concave. Along the optimal path, the consumption per capita $C_t | t \geq 0$ evolves in a stochastic way.
In the constellation of investment opportunities existing in the economy, consider a marginal project that yields a cost $I_0$ today and generates a single benefit $B_t$ at date $t$, where $B_t$ is potentially uncertain and statistically related to stochastic process governing aggregate consumption. At the margin, investing in this project raises the discounted expected utility of the representative agent by

$$\Delta V = -I_0 u'(C_0) + e^{-\rho t} E[B_t u'(C_t)] = u'(C_0) \times NPV,$$

(1)

with

$$NPV = -I_0 + e^{-r_t} E[B_t]$$

(2)

and

$$e^{-r_t} = e^{-\rho t} \frac{E[B_t u'(C_t)]}{u'(C_0) E[B_t]}.$$  

(3)

This means that the increase in the representative agent’s intertemporal welfare generated by the project is proportional to its Net Present Value (NPV) when using the appropriate risk-adjusted discount rate $r_t$ to discount the project’s expected future benefit. This supports the use of the NPV rule to evaluate the investment project. The efficient discount rate defined by equation (3) depends upon the risk profile of the project and its maturity. Notice also that along the optimal path, all socially desirable investments must have been implemented so that $NPV = 0$ is an equilibrium condition, yielding the property that $e^{r_t}$ be equal to $EB/I_0$, which is the expected gross rate of return of the project. In other words, the socially desirable discount rate of an asset must also be its expected rate of return at equilibrium. Because entrepreneurs implementing the project must compensate stakeholders by offering this return in expectation, this induces them to invest in it only if its expected return is larger than $r_t$. This is equivalent to using $r_t$ as the discount rate to evaluate the project. This is an illustration of the first theorem of welfare economics. Consider first the case in which the future benefit is certain, or more generally when it is independent of $C_t$. This yields the risk-free discount rate $r_{ft}$, i.e., the interest rate, which is defined as follows:

$$\exp(-r_{ft} t) = \exp(-\rho t) \frac{E[u'(C_t)]}{u'(C_0)}.$$  

(4)

Consider alternatively an investment consisting in abatement frontloading: One increases the abatement effort today by one ton of carbon dioxide. This allows for abating $e^{-\delta t}$ less tons of carbon dioxide in $t$ years, where $\delta$ is the rate of decay of carbon dioxide. This implies that the concentration of CO$_2$ is unaffected by this intertemporal reallocation at any time after date $t$. Because the initially optimal allocation satisfies the carbon budget constraint, this new allocation does also satisfy this constraint. Let $A'_t|_{t \geq 0}$ denote the dynamics of marginal abatement costs along the optimal allocation of climate efforts. This investment yields an initial cost $I_0 = A'_0$ and generates a future benefit $B_t = \exp(-\delta t) A'_t$. Therefore, it is socially desirable that this benefit be discounted at rate $g_t$ satisfying the following condition:

$$\exp(-g_t t) = \exp(-(\delta + \rho) t) \frac{E[A'_t u'(C_t)]}{u'(C_0) E[A'_t]}.$$  

(5)

If the climate policy is decentralized through a market for emission permits, marginal abatement costs will be equalized across firms and individuals, and will be equal to the equilibrium
carbon price $p_t$. Remember now that $g_t$ can also be interpreted as the equilibrium expected return: $\exp(g_t t)$ must be equal to $E A_t' / A_0'$, i.e., to $E p_t / p_0$. This means that $g_t$ is the efficient growth rate of expected carbon price. Combining equations (4), (5) and $\exp(g_t t) = E p_t / p_0$ yields

$$g_t = \frac{1}{t} \log \left( \frac{E p_t}{p_0} \right) = r_{ft} + \delta + \frac{1}{t} \log \left( \frac{E[A_t'] E[u'(C_t)]}{E[A_t' u'(C_t)]} \right).$$

(6)

The left-hand side of this equality is the annualized growth rate of expected carbon price between dates 0 and $t$. Suppose first that $A_t'$ is constant, or more generally, statistically independent of $C_t$. In that case, the last term in the right side of this equality vanishes. This implies that the efficient growth rate of (expected) carbon price must be equal to the sum of the interest rate and the rate of natural decay of carbon dioxide in the atmosphere. This is the well-known Hotelling’s rule adapted to carbon pricing under a fixed intertemporal carbon budget (Schubert (2008)). More generally, equation (6) tells us that $g_t$ is larger or smaller than $r_{ft} + \delta$ depending upon whether the last term in the right-hand side of this equality is positive or negative. The following proposition is a direct consequence of the covariance rule (Gollier (2001), Proposition 15).7

**Proposition 1.** The growth rate of the expected carbon price that supports the optimal temporal allocation of abatement efforts is larger (smaller) than the sum of the interest rate and the rate of decay of carbon dioxide if the marginal abatement cost and aggregate consumption are (anti-)comonotone.

From the social point of view, facing a positive correlation between marginal abatement costs and aggregate consumption is good news. It means that the worst-case scenarios in terms of abatement costs arise when aggregate consumption is large, i.e., when the marginal abatement effort has a smaller utility impact. This means that abating more in the future reduces the macroeconomic risk. It raises the collective willingness to postpone abatement efforts. It reduces the efficient carbon price today, in exchange for a larger growth rate of this price in the future to satisfy the intertemporal carbon budget. From the individual point of view, investors who implement the abatement frontloading must be compensated for the fact that the benefit of doing so has a positive beta, in the sense that the return of this investment is smaller when other assets also perform poorly in the economy. Because the return of abatement frontloading is the growth rate of carbon price, this compensation takes the form of a growth rate of expected carbon price larger than the sum of the interest rate and the rate of natural decay.

Let us now consider the following special case. Suppose that relative risk aversion is a constant $\gamma$. Suppose also that aggregate consumption and marginal abatement costs evolve according to the following stochastic process:

$$dc_t = \mu_c dt + \sigma_c dz_t$$

$$da_t' = \mu_p dt + \phi \sigma_c dz_t + \sigma_w dw_t,$$

with $c_t = \log C_t$ and $a_t' = \log A_t'$, and where $z_t$ and $w_t$ are two independent standard Wiener processes.8 This means that the logarithm of aggregate consumption and marginal costs

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7Two random variables $(X, Y)$ are said to be comonotone iff for any pair $(s, s')$ of states of nature, $(X(s) - X(s'))(Y(s) - Y(s'))$ is non-negative. Anti-comonotonicity is defined symmetrically.

8Without loss of generality, I normalize $C_0$ and $A_0'$ to unity.
are jointly normally distributed. Parameters $\mu_c$ and $\sigma_c$ are respectively the trend and the volatility of consumption growth. The trend of growth of the marginal abatement cost, and thus of the carbon price, is given by parameter $\mu_p$. The volatility of the marginal abatement cost has an independent component $\sigma_w$ and a component coming from its correlation with economic growth. Notice that $\phi$ can be interpreted as the elasticity of marginal abatement costs to unanticipated changes in aggregate consumption.

We provide a formal proof of the following proposition in the Appendix, together with the characterization of the risk-free rate and the aggregate risk premium. It is an application of the Consumption-based CAPM, and to its illustration to the pricing a non-renewable resource (Gaudet and Khadr (1991)).

**Proposition 2.** Suppose that relative risk aversion is constant and that the logarithms of aggregate consumption and marginal abatement costs follow a bivariate Brownian process. Then, the growth rate of the expected carbon price that supports the optimal temporal allocation of abatement efforts must be equal to the sum of three terms:

- $\delta$: the rate of natural decay of greenhouse gas in the atmosphere;
- $r_f$: the interest rate in the economy;
- $\phi\pi$: the abatement risk premium, which is the product of the income-elasticity ($\phi$) of marginal abatement cost by the aggregate risk premium ($\pi$) in the economy.

In short, we have that

$$ g = \delta + r_f + \phi\pi, $$

(9)

where the interest rate $r_f$ and the aggregate risk premium $\pi$ are characterized in the Appendix. This result tells us that the CCAPM risk premium for carbon permits holds with a CCAPM “carbon beta” being equal to the income-elasticity $\phi$ of the marginal abatement cost. An immediate consequence of Proposition 2 is that the growth rate of expected carbon price is larger (smaller) than the sum of the interest rate and the rate of decay of carbon dioxide if the income-elasticity of marginal abatement costs is positive (negative). This is a special case of Proposition 1.

Under the stochastic process (7)-(8), the estimation of the key parameter $\phi$ is rather simple. Indeed, this system implies that

$$ \Delta \log(A'_t) = a + \phi \Delta \log(C_t) + \varepsilon_t, $$

(10)

where $\Delta \log(A'_t)$ and $\Delta \log(C_t)$ are respectively changes in log marginal cost and in log consumption, and $\varepsilon_t$ is an independent noise that is normally distributed. This means that, under these assumptions, the OLS estimator of the slope of this linear equation is an unbiased estimator of the income-elasticity of the marginal abatement cost that must be used to determine the efficient growth rate of expected carbon price.

### 3 The determinants of the income-elasticity of marginal abatement costs in a simple two-period model

Proposition 2 provides a simple characterization of the efficient growth rate of expected carbon price that relies on the income-elasticity of the marginal abatement cost. In this
section, we explore the determinants of this income-elasticity. Because the current and future marginal abatement costs depend upon which intertemporal abatement strategy is used, this characterization requires solving the intertemporal carbon allocation problem. This cannot be easily done in a continuous-time framework. In this section, we solve this problem in a simple two-period framework. Suppose that the carbon budget constraint covers only two periods, $t = 0$ and $1$. The production of the consumption good is denoted $Y_0$ and $Y_1$ for periods $1$ and $2$ respectively, where $Y_1$ is uncertain in period $0$. The carbon intensity of the economy in the business-as-usual scenario in period $t$ is denoted $Q_t \geq 0$, so that $Q_t Y_t$ tons of carbon dioxide are emitted in period $t$ under this scenario. The country is committed not to exceed a total emission target $T$ for the two periods. As stated for example in the Paris Agreement, the long-term carbon budget allocated to the countries could be modified depending upon new scientific information about the intensity of the climate change problem for example. In our model, this means that, in period $0$, there may be some uncertainty about what the intertemporal carbon budget $T$ will be in the future.

Compared to the business-as-usual scenario, the country must choose how much to abate in each period. Let $K_t$ denote the number of tons of carbon dioxide abated due to actions implemented in period $t$, so that one can write the carbon budget constraint as follows:

$$e^{-\delta} (Q_0 Y_0 - K_0) + Q_1 Y_1 - K_1 \leq T,$$

where $\delta$ is the rate of natural decay of carbon dioxide in the atmosphere. We hereafter assume that this carbon budget constraint is always binding, so that we can rewrite the abatement in period $1$ as a function of the other variables:

$$K_1 = K_1(K_0, Y_1, T) = e^{-\delta} (Q_0 Y_0 - K_0) + Q_1 Y_1 - T.$$

Because $Y_1$ and $T$ are potentially uncertain, so is the abatement effort $K_1$ in period $1$ that will be necessary to satisfy the intertemporal carbon budget constraint.

Abating is costly. Let $A_0(K_0)$ and $A_1(K_1, \theta)$ denote the abatement cost function in periods $0$ and $1$ respectively. We assume that $A_t$ is an increasing and convex function of $K_t$. In order to allow for technological uncertainty, $A_1$ is a function of parameter $\theta$, which is unknown in period $0$. Consumption in period $t$ is $C_t = Y_t - A_t$.

The problem of the social planner is thus to select the abatement strategy $(K_0, K_1)$ to maximize the intertemporal welfare function subject to the carbon budget constraint:

$$\max_{K_0, K_1} H(K_0, K_1) = u(Y_0 - A_0) + e^{-\rho} E[u(Y_1 - A_1)] \quad s.t. \quad (12).$$

The first-order condition of this problem is written as follows:

$$A_0'(Y_0 - A_0) = e^{-\rho-\delta} E \left[ A_1'(Y_1 - A_1) \right],$$

where $A_t'$ denote the partial derivative of the total abatement cost function with respect to abatement $K_t$.

We know from Proposition 1 that the growth rate of the expected carbon price is larger (smaller) than the interest rate when the marginal abatement cost and aggregate consumption...\footnote{This definition allows us to include in the analysis long-term green investments made in period $0$ that also reduce emissions in period $1$ at zero marginal cost. Under this interpretation, $K_1$ must be interpreted as the abatement in period $1$ net of the abatement generated by investments made in the previous period.}
are (anti-)comonotone. In the remainder of this section, we examine various special cases that highlight some of the factors that determine whether the growth rate of expected carbon price should be larger or smaller than the interest rate plus the rate of natural decay of carbon dixiode. Suppose first that the only source of uncertainty in the economy is related to the exogenous growth of production $Y_1$. In particular, this means that $\theta$ is certain, i.e., there is no uncertainty about the green technological progress. It also means that there is no uncertainty about the intertemporal carbon budget allocated to the country. The only source of correlation between $A'_1$ and $C_1$ comes from the fact that both random variables covary with $Y_1$. In that case, we have that

$$\frac{\partial A'_1}{\partial Y_1} = Q_1 A''_1(K_1, \theta),$$

which is positive. We also have that

$$\frac{\partial C_1}{\partial Y_1} = 1 - Q_1 A'_1(K_1, \theta).$$

We hereafter assume that $Q_1 A'_1$ is smaller than unity. Although it is restrictive, this condition is intuitive, since it means that more production growth cannot be bad news, in spite of the increased abatement effort necessary to compensate the extra emission generated by this production. $Q_1 A'_1$ is the increased abatement cost necessary to compensate for the increased production growth in the business-as-usual scenario. This condition states that production growth always increases consumption, even after taking account of the increased abatement effort to compensate for it under the intertemporal carbon constraint. Thus, under this condition, the marginal abatement cost and aggregate consumption are comonotone. Using Proposition 1, this demonstrates the following proposition.

**Proposition 3.** Suppose that the growth of aggregate production $Y_1$ is the only source of uncertainty in the economy, and that $Q_1 A'_1$ is smaller than unity. Then, it is socially desirable that the growth rate of expected carbon price be larger than the sum of the interest rate and the rate of decay of CO$_2$.

A similar exercise can be done in a context where the only source of uncertainty is related to the intertemporal budget constraint $T$. In that case, a larger budget $T$ implies a smaller abatement effort, and thus a larger share of production available for consumption rather than for abatement efforts. At the same time, because of the convexity of the cost function, the marginal abatement cost is smaller. Thus, aggregate consumption and marginal abatement cost are anti-comonone. This yields the following result.

**Proposition 4.** Suppose that the intertemporal carbon budget $T$ is the only source of uncertainty in the economy. Then, it is socially desirable that the growth rate of expected carbon price be smaller than the sum of the interest rate and the rate of decay of CO$_2$.

Suppose finally that the only source of uncertainty is about $\theta$, which is related to the speed of green technological progress. Suppose that an increase in $\theta$ implies a reduction in both the total and the marginal abatement costs, i.e., that for all $(K_1, \theta)$,

$$\frac{\partial A_1(K_1, \theta)}{\partial \theta} \leq 0 \quad \text{and} \quad \frac{\partial^2 A_1(K_1, \theta)}{\partial K_1 \partial \theta} \leq 0.$$

(15)
A possible illustration is when marginal abatement cost is an uncertain constant, i.e., when $A_1(K_1, \theta)$ is equal to $\alpha + g(\theta)K_1$ with $g' \leq 0$, a case examined by Baumstark and Gollier (2010). In that case, a small $\theta$ means at the same time a large marginal abatement cost and a large total abatement cost, and thus a low aggregate consumption. Thus, $A'_1$ and $C_1$ are anti-comonotone, thereby demonstrating the following proposition.

**Proposition 5.** Suppose that the speed of green technological progress $\theta$ is uncertain. If total and marginal abatement costs are comonotone (condition (15)), it is socially desirable that the growth rate of expected carbon price be smaller than the sum of the interest rate and the rate of decay of CO$_2$.

Up to this point, we only characterized the impact of uncertainty on the optimal growth rate of the carbon price. A more complete analysis would be to characterize its effect on the optimal abatement effort in the first period. This is a more difficult question. In order to address it, let us simplify the problem by assuming that the marginal abatement cost in period 1 is constant but potentially uncertain: $A_1(K_1, \theta) = \theta K_1$. In that case, aggregate consumption in period 1 equals

$$C_1 = Y_1 - \theta \left( e^{-\delta} (Q_0 Y_0 - K_0) + Q_1 Y_1 - T \right).$$

Observe that in that case, the first period abatement $K_0$ has a role similar to saving in the standard consumption-saving problem. Each ton of CO$_2$ "saved" in the first period generates an increase in consumption by $R = \exp(-\delta) \theta$ in the second period, where $R$ can be interpreted as the rate of return of savings. Suppose first that $\theta$ is certain. It is well-known in that case that the uncertainty affecting future incomes raises optimal (precautionary) saving if and only if the individual is prudent (Drèze and Modigliani (1972), Leland (1968), Kimball (1990)).

Applying this result to our context directly yields the following proposition. Notice that because the marginal abatement cost is certain, it must grow at the interest rate in this case.

**Proposition 6.** Suppose that $A_1(K_1, \theta) = \theta K_1$ and that the marginal abatement cost $\theta$ is a known constant. Increasing risk on future production $Y_1$ or on the intertemporal carbon budget $T$ increases the initial abatement effort $K_0$ if and only if the representative agent is prudent.

When the marginal abatement cost is uncertain, the future return of abating more today becomes uncertain in that case. By risk aversion, this reinforces the willingness to abate in the first period because it also reduces the risk borne in the second period. because of this second effect, prudence is sufficient but not necessary in this case.

**Proposition 7.** Suppose that $A_1(K_1, \theta) = \theta K_1$ and that the marginal abatement cost $\theta$ is the only source of uncertainty. Increasing the risk affecting $\theta$ raises the initial abatement effort $K_0$ if the representative agent is prudent.

Proof: Consider two random variables, $\theta_1$ and $\theta_2$, where $\theta_2$ is riskier than $\theta_1$ in the sense of Rothschild and Stiglitz (1970). Let $G_i(K_0) = H_i(K_0, K_1(K_0, Y_1, T))$ denote the corresponding objective function, as described by (13). Let $K_{0i}$ denote the optimal initial

\[^{10}\text{An individual is prudent if and only if the third derivative of } u \text{ is positive.}\]
abatement under distribution \( \theta \) of the marginal abatement cost. The optimal abatement effort \( K_{01} \) under the initial uncertainty \( \theta_1 \) satisfies the first-order condition

\[
A'(K_{01})u'(Y_0 - A_0(K_{01})) = \beta E[\theta_1 u'(Y_1 - \theta_1 K_{11})],
\]

(16)

where \( K_{11} \) is the optimal abatement effort in period 1 under the initial risk \( \theta_1 \), i.e., \( K_{11} = K_1(K_{01}, Y_1, T) \). Because \( G_2 \) is concave in \( K_0 \), we obtain that \( K_{02} \) is larger than \( K_{01} \) if and only if \( G_2'(K_{01}) \) is positive. Using condition (16), this condition can be written as follows:

\[
E[\theta_2 u'(Y_1 - \theta_2 K_{11})] \geq E[\theta_1 u'(Y_1 - \theta_1 K_{11})].
\]

(17)

This is true for any Rothschild-Stiglitz risk increase if and only if function \( v \) is convex, where \( v(\theta) = \theta u'(Y_1 - \theta K_{11}) \) for all \( \theta \) in the joint support of \( \theta_1 \) and \( \theta_2 \). It is easy to check that

\[
v''(\theta) = -2K_{11}u''(Y_1 - \theta K_{11}) + \theta K_{11}^2 u'''(Y_1 - \theta K_{11}).
\]

(18)

Because \( K_{11} \) is positive and \( u'' \) is negative, we see that \( v \) is convex when \( u''' \) is positive.

4 Calibration

In this section, we calibrate the two-period model described in the previous section. A standard approach to climate policy in the western world is based on the hypothesis that the energy transition should be performed within the next 3 decades in order to remain below the 1.5\(^\circ\)C objective with probability \( 1/2 \). We follow for example Metcalf (2018) to decompose the next 3 decades into two periods of 15 years, 2021-2035 and 2036-2050. We examine the case of the European Union (EU-28). We hereafter describe the calibration of this model. We assume a rate of pure preference for the present equaling \( \rho = 0.5\% \) per year, and a constant relative risk aversion of \( \gamma = 3 \).

4.1 Economic growth

The current annual GDP of EU-28 is around 19,000 billions US$ (GUS$). Assuming an annual growth rate of 1.4% per year over the period 2021-2035 yields a total production for this first period estimated at \( Y_0 = 315,000 \) billions US$. The production \( Y_1 \) of the second period is uncertain. A key element of this paper is that the recommended returns of green investments are compatible with the equilibrium returns of other assets in the economy, and with intertemporal social welfare. However, as is well-known, the CCAPM model that we use in this paper has been unable to predict observed asset prices when beliefs are normally distributed as assumed in Section 2. This model yields an interest rate that is too large and an aggregate risk premium that is too low.\(^{11}\) In most of this paper, we use the resolution of these asset pricing puzzles that has been proposed by Barro (2006), who recognized the possibility of infrequent large recessions that are not well represented in U.S. growth data. We follow the calibration proposed by Martin (2013). The production in the first period of 15 years is normalized to unity. The change in log production during the second subperiod

\(^{11}\)See for example Kocherlakota (1996) and Cochrane (2017).
is equal to the sum of 15 independent draws of an annual growth rate $x_i$ whose distribution compounds two normally distributed random variables:

$$\log \left( \frac{Y_1}{Y_0} \right) = \sum_{i=1}^{15} x_i$$  \hspace{1cm} (19)

$$x_i \sim (h_{bau}, 1 - p; h_{cat}, p)$$  \hspace{1cm} (20)

$$h_{bau} \sim N(\mu_{bau}, \sigma_{bau}^2)$$  \hspace{1cm} (21)

$$h_{cat} \sim N(\mu_{cat}, \sigma_{cat}^2).$$  \hspace{1cm} (22)

With probability $1 - p$, the annual growth rate is drawn from a "business-as-usual" normal distribution with mean $\mu_{bau} = 2\%$ and volatility $\sigma_{bau}^2 = 2\%$. But with a small probability $p = 1.7\%$, the annual growth rate is drawn from a "catastrophic" normal distribution with a large negative $\mu_{cat} = -35\%$ and a large volatility $\sigma_{cat}^2 = 25\%$. In Table 1, we describe the value of the parameters of the model that are used as a benchmark. The order of magnitude of the parameters of the production growth process is in the range of what has been considered by Barro (2006) and Martin (2013). It yields an annual trend of growth of 1.37\% and an annual volatility of 6.12\%. It also generates an expected production of $Y_1 = 387,000$ billions USD (GUS$) in the second period.

### 4.2 Emissions, decarbonization and decay

The EU-28 currently emits 4.4 GtCO$_2$e per year. Under the Business-As-Usual (BAU), we assume that this flow is maintained over each of the 15 years of the first period, implying 66 GtCO$_2$e emitted in this scenario. When compared to the production $Y_0$ estimated above, this yields a carbon intensity of $Q_0 = 2.10 \times 10^{-4}$ GtCO$_2$e/GUS$. Even without any mitigation policy, the world economy have benefitted from a natural reduction of the energy intensity of its global production over the recent decades. According to Clarke et al. (2014), the average of decline of the energy intensity has been approximately 0.8\% per year between 1970 and 2010. This is why we assume in this calibration exercise that the carbon intensity in the second period goes down to $Q_1 = 1.85 \times 10^{-4}$ GtCO$_2$e/GUS$ in the BAU. This implies an expected total emission of around 72 GtCO$_2$e in the second period under the BAU.

There exists an intense debate about the half-life of carbon dioxide in the atmosphere, and thus on its rate of natural decay. It appears that the carbon cycle is highly non linear, and involves complex interactions between the atmosphere and different layers of the oceans. The existing literature on the half-life of carbon dioxide offers a wide range of estimates, from a few years to several centuries. We conservatively assume a rate of natural decay of CO$_2$ in the atmosphere of 0.5\% per year. This implies a total expected emission net of the natural decay for the European Union over the period 2021-2050 in the BAU around 133 GtCO$_2$e.

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12It is interesting to compare the long run risk generated in this model to the one examined by Nordhaus (2018) and Christensen et al. (2018). They uses a survey of a panel of experts to characterize the uncertainty in estimates of global output for the period 2010-2050. Experts were requested to estimate the average annual growth rate of the period. The resulting estimates were best fit using a normal distribution, with a mean of 2.50\% and a standard deviation of 1.13\%. This yields a standard deviation of $\log(Y_{2050}/Y_{2010})$ equaling $40 \times 1.13\% = 45.2\%$. This should be compared to the standard deviation of $\sqrt{40 \times 6.12\%} = 38.7\%$ for this variable in our model. Thus, we assume long run output uncertainty whose intensity is similar to Nordhaus (2018).

13For a survey on this matter, see Archer et al. (2009).
<table>
<thead>
<tr>
<th>parameter</th>
<th>value</th>
<th>description</th>
</tr>
</thead>
<tbody>
<tr>
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<td>0.5%</td>
<td>annual rate of pure preference for the present</td>
</tr>
<tr>
<td>$\gamma$</td>
<td>3</td>
<td>concavity of utility function</td>
</tr>
<tr>
<td>$p$</td>
<td>1.7%</td>
<td>annual probability of a macroeconomic catastrophe</td>
</tr>
<tr>
<td>$\mu_{bau}$</td>
<td>2%</td>
<td>mean growth rate of production in a business-as-usual year</td>
</tr>
<tr>
<td>$\sigma_{bau}$</td>
<td>2%</td>
<td>volatility of the growth rate of production in a business-as-usual year</td>
</tr>
<tr>
<td>$\mu_{cat}$</td>
<td>-35%</td>
<td>mean growth rate of production in a catastrophic year</td>
</tr>
<tr>
<td>$\sigma_{cat}$</td>
<td>25%</td>
<td>volatility of the growth rate of production in a catastrophic year</td>
</tr>
<tr>
<td>$Y_0$</td>
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<td>production in the first period (in GUS$)</td>
</tr>
<tr>
<td>$\delta$</td>
<td>0.5%</td>
<td>annual rate of natural decay of CO$_2$ in the atmosphere</td>
</tr>
<tr>
<td>$Q_0$</td>
<td>$2.10 \times 10^{-4}$</td>
<td>carbon intensity of production in period 0 (in GtCO$_2$/GUS$)</td>
</tr>
<tr>
<td>$Q_1$</td>
<td>$1.85 \times 10^{-4}$</td>
<td>carbon intensity of production in period 0 (in GtCO$_2$/GUS$)</td>
</tr>
<tr>
<td>$\mu_T$</td>
<td>40</td>
<td>expected carbon budget (in GtCO$_2$)</td>
</tr>
<tr>
<td>$\sigma_T$</td>
<td>10</td>
<td>standard deviation of the carbon budget (in GtCO$_2$)</td>
</tr>
<tr>
<td>$b$</td>
<td>1.67</td>
<td>slope of the marginal abatement cost functions (in GUS$/GtCO_2$)</td>
</tr>
<tr>
<td>$a_0$</td>
<td>23</td>
<td>marginal cost of abatement in the BAU, first period (in GUS$/GtCO_2$)</td>
</tr>
<tr>
<td>$\mu_{\theta}$</td>
<td>2.31</td>
<td>expected future log marginal abatement cost in BAU</td>
</tr>
<tr>
<td>$\sigma_{\theta}$</td>
<td>1.21</td>
<td>standard deviation of future log marginal abatement cost in BAU</td>
</tr>
</tbody>
</table>

Table 1: Benchmark calibration of the two-period model.

### 4.3 Carbon budget

In the most recent report of the IPCC (IPCC (2018)), the goal of not exceeding a 1.5°C increase in temperature compared to the pre-industrial age is estimated to be compatible with a median carbon budget of 770 GtCO$_2$ in early 2018. Given that we have emitted around 40 Gt of greenhouse gases per year since then, we assume that this global carbon budget has now been reduced to 730 Gt. There is a debate about how to share this total carbon budget among the different countries. Let us take the conservative (and ethically sounded) approach of sharing the budget on a per capita basis. Because the European Union is home for roughly 7% of the world population, we assume that EU-28 should be allocated a carbon budget of approximately 50 GtCO$_2$. Let us further assume that four-fifth of this budget could be consumed between 2021 and 2050. This gives an expected carbon budget for EU-28 for that period equalling $\mu_T=40$ GtCO$_2$. Compared to the global emission of 133 GtCO$_2$, this represents a global abatement effort of 93 GtCO$_2$, or a reduction of more than 70% of the global BAU emissions in the EU-28 during the next 3 decades.

There is of course much uncertainty about what will be the actual carbon budget that will emerge from the international negotiations in the next 3 decades, and from the resolution of the uncertainty about the intensity of climate change. We model this uncertainty by assuming that $T$ is normally distributed with mean $\mu_T$ and standard deviation $\sigma_T = 10$ GtCO$_2$. 

15
4.4 Abatement costs

We assume that the abatement cost function is quadratic:

\[ A_t(K_t) = a_t K_t + \frac{1}{2} b K_t^2. \]  \hspace{1cm} (23)

An important element of our model is related to how the marginal abatement cost (MAC) changes with the ambition of the mitigation policy. The answer to this question is given by the MAC slope coefficient \( b \), which tells us by how much the marginal abatement cost increases when the abatement effort increases by 1 Gt of CO\(_2\)e. The researchers behind the MIT Emissions Prediction and Policy Analysis (EPPA, Morris et al. (2012)) have developed computable general equilibrium models with a very detailed energy sector. They have estimated the shadow price of carbon associated to various carbon budgets for different regions of the world, thereby generating regions-specific MAC curves. We used their analysis of the MAC curve for the European Union in 2020 to estimate that the MAC increases by 25 US$ whenever the annual abatement effort is increased by 1 GtCO\(_2\)e. Expressed for a period of 15 years, this suggests \( b = 1.67 \text{ GUS$}/\text{GtCO}_2\text{e}^2 \). We assume that \( b \) is certain and constant over time.

Parameter \( a_t \) measures the MAC along the BAU scenario. For the first period, we estimate it by price of carbon permit in the summer of 2018 on the EU-ETS market, around 23 GUS$/GtCO_2e. The full elimination of the 66 GtCO\(_2\)e emitted in the first period would cost around 5,000 GUS$, or 1.6% of GDP in the first period.

The MAC in the BAU during the second period is uncertain. Anticipating green innovations would suggest using \( a_1 \) smaller than \( a_0 \), at least in expectation. By how much smaller remains an open question. In order to estimate the degree of uncertainty that surrounds abatement costs in the second period of our analysis, we have used a set of AIM models scrutinized by the Working Group III for the Fifth Report of the IPCC (Clarke et al. (2014)). In the associated database\(^{14}\), we have collected the 374 estimations of carbon prices for 2030 that are in line with the objective of not exceeding 450ppm over the century. These estimates differ by the IAM model used for the estimation, and by the assumed technological progresses available at that time horizon. We depict the histogram of these MAC estimates for 2030 in Figure 2. The distribution of these estimates is heavily skewed to the right, which suggests using a lognormal distribution for \( a_1 = \theta \). The standard deviation of the log MAC in this sample is equal to \( \sigma_\theta = 1.21.^{15} \) The standard deviation of the future MAC at the BAU is equal to 38 US$/tCO_2e, which is in the range of the MAC uncertainty measured by Kuik et al. (2009) for a time horizon of 15 years.\(^{16}\)

The trend of reduction in the MAC in our calibration is aligned with the assumption made by Nordhaus (2018) that the cost of the backstop technology declines at a rate of 0.5% per year. In our benchmark calibration, we assume that \( \log(\theta) \) is normally distributed with mean \( \mu_\theta = 2.31 \) and standard deviation \( \sigma_\theta = 1.21 \). This yields an expected MAC in the BAU around 21GUS$/GtCO_2e. The 8% reduction in the expected MAC under the BAU measures the green innovations that are expected to emerge in the next 15 years.

\(^{14}\)https://tntcat.iiasa.ac.at/AR5DB

\(^{15}\)Because these estimates are based on an ambitious abatement target, the mean value of the carbon price in this sample is not useful for the estimation of the expected MAC in the BAU.

\(^{16}\)These authors performed a meta-analysis of MAC estimates in the literature, and observed a standard deviation of MAC of 27.9 and 52.9 euros per tCO\(_2\)e respectively for 2025 and 2050.
### Table 2: Description of the optimal solution in the benchmark case.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Value</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>$K_0$</td>
<td>31.10</td>
<td>optimal abatement in the first period (in GtCO₂e)</td>
</tr>
<tr>
<td>$E[K_1]$</td>
<td>66.30</td>
<td>optimal expected abatement in the second period (in GtCO₂e)</td>
</tr>
<tr>
<td>$p_0$</td>
<td>74.90</td>
<td>optimal carbon price in the first period (in US$/tCO₂e)</td>
</tr>
<tr>
<td>$E[p_1]$</td>
<td>132.00</td>
<td>optimal expected carbon price in the second period (in US$/tCO₂e)</td>
</tr>
<tr>
<td>$g$</td>
<td>3.76</td>
<td>annualized growth rate of expected carbon price (in %)</td>
</tr>
<tr>
<td>$r_f$</td>
<td>0.98</td>
<td>annualized interest rate (in %)</td>
</tr>
<tr>
<td>$\pi$</td>
<td>2.51</td>
<td>annualized systematic risk premium (in %)</td>
</tr>
<tr>
<td>$\phi$</td>
<td>1.00</td>
<td>OLS estimation of the income-elasticity of the marginal abatement cost</td>
</tr>
</tbody>
</table>

4.5 Results

We solved the first-order condition (14) numerically by using the Monte-Carlo method. We draw 350,000 random triplets $(Y_1, \theta, T)$ that we approximated the expectation of the right-hand side of this equality by an equally weighted sum of this random sample. In Table 2, we describe the optimal solution of this problem under the calibration of the parameters described in Table 1. We obtain equilibrium asset prices that are in line with the observed real interest rate of 1% and systematic risk premium of 2% that have been observed in the United States during the last century (Kocherlakota (1996)). The expected optimal abatement is much larger in the second period than in the first one. This is partly due to the anticipation of a larger price of carbon in the second period. In expectation, the annualized growth rate of the carbon price equals 3.76%. This is much larger than the sum of the natural rate of decay of CO₂ and the interest rate, which is equal to 1.69%. This is due to the fact that at the optimum, the marginal abatement cost is positively correlated with aggregate consumption, as shown in Figure 3. In fact, the OLS estimation of the income-elasticity of the marginal abatement cost is $\phi \simeq 1.00$.17

As observed by Metcalf (2018), Aldy (2017) and Hafstead et al. (2017), carbon price predictability is the most important feature of a climate policy for the business community as it plans long-term investments in line with the energy transition. For example, Metcalf (2018) proposes to fix the annual growth rate of carbon price at 4% (plus inflation) as long as the path of emissions is in line with the objective. However, under uncertainty, the efficient growth rate of carbon price must be uncertain in this model because the resolution of the uncertainty affecting economic growth, green innovations and the carbon budget needs to be translated into a variable carbon price in the second period. We represented the distribution of the carbon price $p_1$ and its annualized growth rate respectively in Figures 4 and 5. The standard deviation of this annualized growth rate is equal to 2.4% per annum. Contrary to the above-mentioned view, it is desirable that this risk be borne by the business community. It reflects the uncertainties associated to the price of carbon necessary to satisfy the intertemporal carbon budget constraint. This constraint translates into an uncertain abatement effort in the second period, as described in Figure 6. Investment decisions in

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17Because consumption and marginal abatement costs are not log normal, equation (33) should not be used to estimate the optimal growth rate of expected carbon price.
energy transition should take account of these uncertainties. The attractiveness of green investments should come from their expected return rather than from their reduced riskiness, something that cannot be guaranteed under a rigid carbon budget. The large uncertainty surrounding future carbon prices is a consequence of the cost-efficiency approach associated to a rigid carbon budget constraint.

The high uncertainty about the second period carbon price is also the consequence of the cost-efficiency approach used in this paper. In the alternative cost-benefit approach, the absence of green innovation would be partially compensated by allowing emissions to grow.\textsuperscript{18} This is not possible if one takes the carbon budget constraint seriously. It implies that the carbon price has to grow faster under the cost-efficiency approach in this adverse scenario. Ex ante, this means that the carbon price uncertainty is larger.

What is the welfare cost of fighting climate change for the next three decades? To address this question, we measure welfare associated to a policy by the constant consumption level that generates the same discounted expected utility generated by that policy. Under the optimal carbon pricing rule, this constant consumption level is equal to 330,020 GUS$. This should be compared to the constant consumption level of 332,560 GUS$ that is obtained with the zero ambition strategy, i.e., when $K_0$ and $K_1$ are zero. This means that fighting climate change has an effect on intertemporal welfare that is equivalent to a permanent reduction of consumption by 0.76%.

4.6 The welfare cost of delaying action

We have seen in the introduction that most calibrations of cost-efficiency IAM models yield a growth rate of carbon price that is much larger than the interest rate. Because these models assume no uncertainty, they imply a suboptimal allocation of the abatement effort over time, with a lack of effort in the short run, and too much effort in the long run. This may be due to the political command imposed to these calibrations. In this section, we are interested in measuring the welfare cost of this inefficiency. Our findings are summarized in Table 3.

If the EU maintains the price of permits at its 2018 level (23 US$/tCO_2) for the next 15 years, it will be forced to increase it to almost 180 US$/tCO_2 in expectation during the second period, which corresponds to an annual growth rate of 13.7%. This vastly inefficient intertemporal allocation of efforts yields a welfare loss that is equivalent to a permanent reduction of consumption by 1.04%. Compared to the efficient policy, this represents an increase in welfare loss by 0.28%, from 0.76%. In short, this means that postponing the effort by 15 years has an effect on welfare which is equivalent to reducing consumption by a quarter during the next three decades, a 37% increase in the welfare cost of fighting climate change compared to the efficient policy.

It is noteworthy that the selection of an initial carbon price of 50 US$/tCO_2, halfway between the BAU and the optimal carbon prices, yields a growth rate of expected carbon price of 7.5% per year, not far from what IAM models suggest. The welfare loss associated to this less inefficient policy is only 8% larger than when using the efficient policy.

\textsuperscript{18}This observation implies that marginal abatement cost and marginal abatement benefits are positively correlated, as explained by Stavins (2019).
<table>
<thead>
<tr>
<th>$p_0$</th>
<th>$E[p_1]$</th>
<th>$g$</th>
<th>$K_0$</th>
<th>$E[K_1]$</th>
<th>Welfare loss</th>
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<tr>
<td>23</td>
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<td>13.71</td>
<td>0.00</td>
<td>95.16</td>
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<td>30</td>
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</tbody>
</table>

Table 3: Cost of delaying the abatement effort. The initial price $p_0$ is arbitrarily selected between the BAU level (23 US$/tCO_2$) and its efficient level (75 US$/tCO_2$). The welfare loss measures the reduction (in %) in the constant welfare-equivalent consumption level compared to the no-abatement strategy.

4.7 Risk sensitivity analysis

Table 4 provides some information about the sensitivity of our results to the intensity of the exogenous risk of the model. The most interesting comparison to the benchmark is obtained when all sources of risk are switched off. Suppose that all $s$ are reduced to zero, together with the probability of catastrophe. To preserve the mean growth rate of output, we reduced the mean growth rate to $\mu_{bau}$ to 1.37%. In this risk-free economy, we know that the efficient growth rate of carbon prices must be equal to the sum of the interest rate and the rate of natural decay. No risk premium should be included. However, the absence of uncertainty switches off the precautionary motive to reduce the interest rate, which goes up to $r_f = 4.43\%$ in this context. This yields an efficient growth rate of carbon price of $r_f + \delta = 4.93\%$. The large discount rate implies that very little effort is made in the first period, with a low initial carbon price. In the benchmark calibration, the reason for why much of the mitigation effort is postponed comes from the fact that the MAC is positively correlated with aggregate consumption. In this alternative context with no risk, there is an even stronger argument for delaying the effort, namely, the absence of any precautionary motive to invest.

In the fourth column entitled "no catastrophe", we have solved the model by using the benchmark calibration except for the probability of catastrophe $p$ that has been switched to zero, combined with a reduction of $\mu_{bau}$ to 1.37% in order to leave $E[Y_1]$ unchanged. This has the effect to raise the interest rate and to reduce the systematic risk premium to unrealistic low levels. This observation justifies our choice of introducing macroeconomic catastrophes à la Barro in our calibration. The consequence of reducing the macro risk is to make the growth rate of expected carbon price smaller than the sum of the interest rate and the rate of natural decay, as suggested by our theoretical results. However, because the systematic risk premium is marginal in the absence of catastrophe, the difference between the two is small, as suggested by equation (33), which is an approximation in this non-gaussian calibration.

In the last two columns of Table 4, we document the results of simulations in which

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19 This precautionary motive to reduce the discount rate is best illustrated in the so-called "extended Ramsey rule" (equation (31) of the Appendix). For more details, see for example Gollier (2016).
Table 4: Risk sensitivity analysis. The "no risk" context is obtained by equalizing all standard deviations to zero, by reducing the probability of catastrophe to zero, and by replacing $\mu_{\text{bau}}$ by 1.37% to preserve the expected growth rate of production as in the benchmark. The "no catastrophe" context is obtained by shifting the probability of catastrophe $p$ to zero, and by reducing the trend of growth to $\mu_{\text{bau}}$ to 1.37%. The "no macro risk" context combines these changes with the shift of the volatility $\sigma_{\text{bau}}$ to zero. In the "no tech risk" context, we switched $\sigma_{\theta}$ to zero compared to the benchmark. In the "no budget risk" case, we reduced $\sigma_{T}$ to zero compared to the benchmark. Units are as in Table 2.

<table>
<thead>
<tr>
<th>variable</th>
<th>benchmark</th>
<th>no risk</th>
<th>no cata.</th>
<th>no macro risk</th>
<th>no tech risk</th>
<th>no budget risk</th>
</tr>
</thead>
<tbody>
<tr>
<td>$K_0$</td>
<td>31.10</td>
<td>21.80</td>
<td>26.20</td>
<td>25.90</td>
<td>27.90</td>
<td>30.90</td>
</tr>
<tr>
<td>$E[K_1]$</td>
<td>66.30</td>
<td>73.10</td>
<td>69.30</td>
<td>69.40</td>
<td>69.30</td>
<td>66.50</td>
</tr>
<tr>
<td>$p_0$</td>
<td>74.90</td>
<td>59.40</td>
<td>66.80</td>
<td>66.30</td>
<td>69.60</td>
<td>74.60</td>
</tr>
<tr>
<td>$E[p_1]$</td>
<td>132.00</td>
<td>124.00</td>
<td>137.00</td>
<td>137.00</td>
<td>126.00</td>
<td>132.00</td>
</tr>
<tr>
<td>$g$</td>
<td>3.76</td>
<td>4.93</td>
<td>4.77</td>
<td>4.83</td>
<td>3.94</td>
<td>3.80</td>
</tr>
<tr>
<td>$r_f$</td>
<td>0.98</td>
<td>4.43</td>
<td>4.22</td>
<td>4.39</td>
<td>1.22</td>
<td>1.23</td>
</tr>
<tr>
<td>$\pi$</td>
<td>2.51</td>
<td>-</td>
<td>0.12</td>
<td>0.00</td>
<td>2.29</td>
<td>2.28</td>
</tr>
<tr>
<td>$\phi$</td>
<td>1.00</td>
<td>-</td>
<td>0.66</td>
<td>-24.10</td>
<td>1.01</td>
<td>0.96</td>
</tr>
</tbody>
</table>

4.8 Parameter sensitivity analysis

We now turn to the sensitivity analysis related to the non-risk parameters of the model. We first double the expected carbon budget from $\mu_{T} = 40$ to 80 GtCO$_2$e. This increases the income-elasticity of the MAC and the efficient growth rate of carbon price. This implies a reduction of the carbon price in the first period by almost 40%. We also examined the effect of increasing the trend of reduction of the MAC in the BAU from around 0.5% to 1% per annum, but that reduces the optimal growth rate of carbon price only by 0.04%. The optimal climate policy is very sensitive to the curvature coefficient $b$ of the abatement function. In line with Proposition 3 and the intuition that supports it, doubling the curvature more than doubles of the income-elasticity of the MAC, which in turn implies an increase in the optimal growth rate of carbon prices.

The last two columns of Table 5 are related to the Nordhaus-Stern controversy on the
Table 5: Parameter sensitivity analysis. In the 'doubling carbon budget' scenario, we increase the expected carbon budget from $\mu_T = 40$ to 80 GtCO$_2$e. In the 'more green innovation' scenario, we double the annual rate of reduction of the MAC in the BAU, so that $\mu_\theta$ is reduced from 2.31 to 2.25. We double the curvature coefficient of the abatement cost function to $b = 3.34$ in the scenario entitled 'doubling cost curvature'. In 'Nordhaus', we increase the rate of pure preference for the present from $\rho = 0.5\%$ to 1.5%, and we reduce relative risk aversion $\gamma$ from 3 to 1.45. Finally, in 'Stern', we reduce $\rho$ to 0.1% and $\gamma$ to 1. Units are as in Table 2.

discount rate. Our benchmark calibration was made compatible with observed asset prices. Following Barro (2006), we introduced macro catastrophes and we assumed a constant relative risk aversion equaling $\gamma = 3$. We also used a rate of pure preference for the present equaling $\rho = 0.5\%$. These two coefficients are subject to an intense debate in our profession. Nordhaus (2018) uses a larger $\rho = 1.5\%$, whereas Stern (2007) uses a smaller $\rho = 0.1\%$. Both use a smaller $\gamma$ of 1.45 for Nordhaus, and 1 for Stern. As illustrated by Table 5, this lower curvature of the utility function implies an equilibrium interest rate which is too large, and a risk premium which is too small. As expected, the Nordhaus’ calibration of our collective preferences yields a much smaller initial carbon price and a too large growth rate of carbon prices compared to the Stern’s calibration.²⁰

4.9 An alternative approach: Epstein-Zin preferences

As explained in section 4.1, we used the Barro’s extreme events argument in this calibration in order to solve the asset pricing puzzles that are inherent to the CCAPM model. In this section, we explore the alternative standard resolution of these puzzles that has been provided by Bansal and Yaron (2004). This 'Long Run Risk' (LRR) model has two key ingredients. First, the representative agent is assumed to be endowed with Epstein-Zin recursive preferences, with a large degree of risk aversion, and a low degree of aversion to consumption fluctuations. Second, the growth rate of production is sensitive to a slow-moving state variable.

To adapt our model to this alternative resolution of the asset pricing puzzles, we perform

²⁰Under a cost-benefit approach, Nordhaus would also assume a larger carbon budget than Stern. This is not taken into account in the discussion based on an exogenous carbon budget.
Table 6: Epstein-Zin calibration of the two-period model. All other parameters are as described in the bottom part of Table 1.

<table>
<thead>
<tr>
<th>Value</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\rho$</td>
<td>0.15% monthly rate of pure preference for the present</td>
</tr>
<tr>
<td>$\gamma$</td>
<td>10 relative aversion to risk</td>
</tr>
<tr>
<td>$\alpha$</td>
<td>2/3 relative aversion to consumption fluctuations</td>
</tr>
<tr>
<td>$\mu$</td>
<td>0.15% unconditional expected monthly growth rate of production</td>
</tr>
<tr>
<td>$\sigma$</td>
<td>0.78% st. dev. of shocks to the monthly growth rate of production</td>
</tr>
<tr>
<td>$\sigma_e$</td>
<td>0.034% st. dev. of shocks to predictable component of the monthly growth rate of production</td>
</tr>
<tr>
<td>$k$</td>
<td>0.979 persistence coefficient of the expected growth rate process</td>
</tr>
</tbody>
</table>

two changes in our model. First, the objective function $H$ now takes the following form:

$$H(K_0, K_1) = u(Y_0 - A_0) + e^{-\rho}u(\psi),$$

where $\psi$ is the certainty equivalent of future consumption:

$$v(\psi) = Ev(Y_1 - A_1).$$

As is standard, we assume that $u$ and $v$ belong to the family of power functions, with $u(c) = c^{1-\gamma}/(1-\gamma)$ and $v(c) = c^{1-\alpha}/(1-\alpha)$. Parameter $\gamma$ can be interpreted as the aversion to consumption fluctuations, which is the inverse of the elasticity of intertemporal substitution. Parameter $\alpha$ is relative risk aversion. The second change concerns the dynamic process governing the per-period growth rate of consumption $x_i$. Technically, we substitute the Barro’s extreme events dynamics (20)-(22) by the following LRR dynamics:

$$x_{i+1} = \mu + z_i + \sigma \varepsilon_{i+1}$$

$$z_{i+1} = kz_i + \sigma \varepsilon e_{i+1}.$$  

The slow-moving state variable $z$ has an AR(1) dynamics with a persistency parameter $k$ close to unity. It is assumed that $\varepsilon$ and $e$ are statistically independent $N(0,1)$ random variables. We calibrate this model using the parameter values used by Bansal and Yaron (2004), who used a monthly frequency. This means that $\log(Y_1/Y_0)$ equals $x_1 + ... + x_{180}$. Because market interest rates have experienced lower values since the publication of this paper, we have reduced the monthly rate of impatience from the 0.2% calibrated by Bansal and Yaron (2004) to 0.15%. The parameter values described in Table 6 replace those presented in the upper part of Table 1.

Because future consumption is not log-normally distributed, there is no analytical solution to the problem of maximizing $H$ under the intertemporal carbon constraint (12). In Table 7, we describe the numerical solution to this problem. The bottom line of this exercise is that the LRR model generates an optimal carbon pricing schedule which does not differ much from our benchmark model. The two models share the same key observation that the CCAPM beta of efficient carbon prices is close to unity. They also share similar levels for the interest rate, the market risk premium and the growth rate of expected carbon prices. It is reassuring
### Table 7: Description of the optimal solution of the long run risk model.

<table>
<thead>
<tr>
<th>variable</th>
<th>value</th>
<th>description</th>
</tr>
</thead>
<tbody>
<tr>
<td>( K_0 )</td>
<td>33.44</td>
<td>optimal abatement in the first period (in Gt( \text{CO}_2 \text{e} ))</td>
</tr>
<tr>
<td>( E[K_1] )</td>
<td>68.79</td>
<td>optimal expected abatement in the second period (in Gt( \text{CO}_2 \text{e} ))</td>
</tr>
<tr>
<td>( p_0 )</td>
<td>78.85</td>
<td>optimal carbon price in the first period (in US$/t\text{CO}_2\text{e})</td>
</tr>
<tr>
<td>( E[p_1] )</td>
<td>135.86</td>
<td>optimal expected carbon price in the second period (in US$/t\text{CO}_2\text{e})</td>
</tr>
<tr>
<td>( g )</td>
<td>3.63</td>
<td>annualized growth rate of expected carbon price (in %)</td>
</tr>
<tr>
<td>( r_f )</td>
<td>0.87</td>
<td>annualized interest rate (in %)</td>
</tr>
<tr>
<td>( \pi )</td>
<td>2.63</td>
<td>annualized systematic risk premium (in %)</td>
</tr>
<tr>
<td>( \phi )</td>
<td>1.00</td>
<td>OLS estimation of the income-elasticity of the marginal abatement cost</td>
</tr>
</tbody>
</table>

to conclude that the main messages of our analysis are independent of the strategy used to solve the classical asset pricing puzzles.

Other researchers have examined assets and carbon pricing in an Epstein-Zin framework. In Bansal et al. (2016), final consumption is also affected by the slow-moving average temperature (and climate damages), a second LRR factor in their model. Although both papers share the objective to explore the role of climate change in asset pricing, the media of the climate impacts are radically different. In Bansal et al. (2016), asset prices are affected by climate damages, whereas they are affected by carbon prices and mitigation efforts in our model. Daniel et al. (2019) also characterize an optimal carbon pricing schedule with Epstein-Zin preferences. Rather than using a cost-efficiency approach as in this paper, they characterize the first-best solution with a climate damage function, considering a much longer time horizon of 300 years decomposed into 7 periods. Climate damages and mitigation costs are uncertain, but they assume no uncertainty about production growth. Because we have shown that the absence of production growth uncertainty implies a negative carbon beta, it should not be a surprise that Daniel et al. (2019) obtain an initial carbon price that is relatively large, with a strongly negative trend of growth.\(^2^1\) In their model contrary to ours, early green tech adoption provides a hedge against the long-term aggregate risk. From our analysis, we conjecture that adding some realistic production growth uncertainty into their model will radically affect their recommendations.

#### 5 Conclusion

The future social and private benefits of most investments in renewable energy are uncertain by nature. This is because these investments are capital-intensive and with very long maturities. One of their crucial social benefits is the reduction in emissions of \( \text{CO}_2 \), whose pricing should be the key driver to induce market players to invest. Under a fixed intertemporal carbon budget constraint, the carbon price sends the right signal about the evolution of both the scarcity of emission permits and the cost of abatement efforts. For the sake of efficiency, it needs to be sensitive to news about the evolution of the residual budget, and about green technological shocks. I have shown in this paper that, along the optimal mitigation path, the

\(^{21}\)In the same spirit to suggest a negative beta, they obtain an initial carbon price that is increasing in the market risk premium.
marginal abatement cost is positively correlated with aggregate consumption. To be more precise, I have shown that the MAC has a CCAPM-beta close to 1. The first consequence of this observation is that the benefits of abatement frontloading should be discounted at a rate larger than the interest rate, in recognition of the fact that it raises the macroeconomic uncertainty. The second consequence is that one should compensate early green entrepreneurs by offering them an expected rate of return that is larger than the risk-free rate. How can we implement this? The answer is simple: Real carbon prices should grow in expectation at a rate larger than the sum of the rate of natural decay of CO$_2$ in the atmosphere and of the interest rate. The risk-adjustment of this growth rate of carbon price should be equal to the aggregate risk premium, which has historically been around 2\% per annum.

The renewable industry has often lobbied to obtain guarantees about future carbon prices, with the claim that it is a necessary condition for a rapid energy transition. They are wrong. Rather than offering guarantees about future prices – a policy which would limit the quality of future price signals, one should offer them a larger expected rate of return for their investments in renewable energies, as a compensation for the risk that these investments yields. Again, this takes the form of planning a larger growth rate of expected carbon prices. Very risk-averse green investors should look for financial products that could hedge the carbon price volatility at market price.
References


——— (2013): “Climate change policy: What do the models tell us?” *Journal of Economic Literature*, 51, 860–872.


Appendix: Proof of Proposition 2

Using equation (3), the risk-adjusted discount rate $\tau_{ct}$ to discount a claim on aggregate consumption must satisfy the following efficiency condition:

$$\exp(-\tau_{ct}t) = \exp(-\rho t) \frac{E[C_tu'(C_t)]}{u'(C_0)E[C_t]}.$$  \hfill (28)

The systematic risk premium $\pi_t$ is the extra expected rate of return of a claim on aggregate consumption over the interest rate that must compensate agents who accept to bear the macroeconomic risk:

$$\pi_t = \tau_{ct} - r_{ft}. \hfill (29)$$

Under the two assumptions of the proposition, combining equation (5) with the property that at equilibrium $\exp(g_t t)$ equals $EA'/A'_0$ implies the following equation:

$$1 = e^{-(\rho + \delta)t} E\left[\frac{A'_t u'(C_t)}{A'_0 u'(C_0)}\right] = e^{-(\rho + \delta)t} E\left[(a'_t - \gamma c_t)\right].$$

Notice that our assumptions implies that $a'_t - \gamma c_t$ is normally distributed with mean $\mu_x - \gamma \mu_c$ and variance $(1 - \gamma \phi)^2 \sigma^2_c + \sigma^2_w$. By Stein’s Lemma, the above condition can then be rewritten as follows:

$$1 = \exp\left(\left(-\rho - \delta + \mu_p - \gamma \mu_c + 0.5(\phi - \gamma)^2 \sigma^2_c + 0.5\sigma^2_w\right) t\right),$$

or, equivalently,

$$\mu_p + 0.5\phi^2 \sigma^2_c + 0.5\sigma^2_w = \delta + \rho + \gamma \mu_c - 0.5\gamma^2 \sigma^2_c + \phi \gamma \sigma^2_c. \hfill (30)$$

In this economy, the following standard CCAPM formula for the risk-free interest rate can be derived from equation (4):

$$r_{ft} = r_f = \gamma + \gamma \mu_c - 0.5\gamma^2 \sigma^2_c. \hfill (31)$$

The systematic risk premium $\pi_t$ is given by equation (29). Using Stein’s Lemma twice to estimate $\tau_{ct}$ given by equation (28) yields the following result:

$$\pi_t = \pi = \gamma \sigma^2_c. \hfill (32)$$

Notice also that, using Stein’s Lemma again, we have that the expected marginal abatement cost satisfies the following condition:

$$E\frac{A'_t}{A'_0} = E \exp(a'_t) = \exp\left(\left(\mu_p + 0.5\phi^2 \sigma^2_c + 0.5\sigma^2_w\right) t\right).$$

This implies that the growth rate $g$ of expected marginal abatement cost is a constant given by

$$g = \frac{dEA'_t/dt}{EA'_t} = \mu_p + 0.5\phi^2 \sigma^2_c + 0.5\sigma^2_w.$$
Because in a decentralized economy, the marginal abatement cost is equal to the price of carbon in all states of nature and at all dates, $g$ can also be interpreted as the growth rate of expected carbon price. Combining these properties implies that one can rewrite condition (30) as follows:

$$g = \delta + r_f + \phi \pi.$$  

(33)

This concludes the proof of Proposition 2. ■
Figure 2: Histogram of the world marginal abatement costs for 2030 extracted from the IPCC database (https://tntcat.iiasa.ac.at/AR5DB). We have selected the 374 estimates of carbon prices (in US$2005/tCO$_2$) in 2030 from the IAM models of the database compatible with a target concentration of 450 ppm.

Figure 3: Monte-Carlo simulation under the benchmark case. For the sake of readability of the figure, we limited the simulation to 50,000 draws of the triplets ($Y_1, \theta, T$) to estimate the optimal abatement strategy. The figure illustrates the positive statistical relation between log consumption growth and the log marginal abatement costs (and thus log carbon price) in the second period. The red curve depicts the OLS estimation in log-log.
Figure 4: Empirical probability distribution of the carbon price \( p_1 \) (in US$\$/tCO\(_2\)e) under the optimal abatement strategy in the benchmark calibration of the two-period model. The Monte-Carlo simulation uses a sample of 350,000 draws of the triplet \((Y_1, \theta, T)\).

Figure 5: Empirical probability distribution of the annualized growth rate of carbon price under the optimal abatement strategy in the benchmark calibration of the two-period model. The growth rate is in percent per year. The mean growth rate is 3.36% and the standard deviation is equal to 2.5%.
Figure 6: Empirical probability distribution of the abatement effort $K_1$ (in GtCO$_2$e) under the optimal abatement strategy in the benchmark calibration of the two-period model.